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Seasonal observations Optimal Stopping of MC State Elimination Algorithm Projections of Markov Models (MCs) Reality and Open Problems

## Optimal Stopping of Seasonal Observations and Calculation of Related Fundamental Matrices

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#### Reality vs Theory ! Resistance levels. Do they exist ?

...despite the upturn in investor sentiment, the Dow met with staunch *resistance* near the 11,700 level yesterday...

...though *support* could materialize just below 1,090 at the SPX's 50-day moving average...

Support for the Dow has materialized in the 11,550 region, with additional *support* from its 10-week trendline hovering just below the average. *Resistance*, meanwhile, should materialize near 10,750 if the bulls remain asleep...

#### **Resistance levels 2**

"Model": Suppose that all believe that

- 1) support level d and resistance level D do exist;
- 2) The stock market is a MC with three states:
- a, bearish = random walk with a negative drift,
- *n*, neutral = random walk,
- b, bullish = random walk with a positive drift.

What is your optimal strategy if stock market is in a, n, b?

Obvious: buy in b, sell in a, wait in n, but only inside of (d, D).

Then, we will have self-fulfilling prophecy - markets will follow the model !

#### **Resistance levels 3**

Suppose that you do not believe but other believe.

What is your optimal strategy if stock market is in a, n, b?

Your strategy will be the same !

Then, again - markets will follow the model !

## Outline

- Motivation done
- Seasonal Observations
- Optimal Stopping (OS) of Markov Chains (MCs)
- State Elimination Algorithm (SEA)
- Projections of MCs
- Key transformation
- Open Problems

### Seasonal observations

The simplest problem of optimal stopping is tossing a die.

The following problem of OS of MC was dubbed previously by an author as OS of "Seasonal" Observations.

A DM (decision maker, player) observes a MC  $(Z_n), Z_n = (U_n, Y_n)$  with two components, where the first one is an "underlying" finite MC  $(U_n)$  with *m* states and given transition matrix *U*, and the other component is one of *m* independent sequences of i.i.d. r.v.  $(Y_n(k))$  with known distr-s  $F = \{f(\cdot|k)\}$ .

There are m dice and which of them is tossed at a given moment is specified by a position of a MC  $(U_n)$ .

The goal of a player is to maximize the discounted expected reward over all possible stopping times. The crucial point is that  $P = U \times F$ . Hidden MC if a DM observes only  $(Y_n)$ .

Theorem (1, Presman, Sonin 2010, Theory of Prob. and Appl.)

(a) There are threshold values  $d_*(1), ..., d_*(m)$ , such that the optimal stopping set  $S_* = \bigcup_j S_*(j), S_*(j) = \{x : g(i,j) \ge d_*(j)\}, j \in B = \{1, ..., m\},$ (b)  $d_* = (d_*(s), s \in B)$  satisfies the equation

 $d_*(s) = \sum_{k \in B} l_*(s,k) \sum_{j \in D_*(k)} g(k,j) f(j|k),$ 

where the matrix  $L_* = \{l_*(s, k), s, k \in B\}$  is defined by the equality

 $L_* = [I - UF_d(D_*)]^{-1}U,$ 

c) only a finite number of steps  $k_*$  is necessary to obtain an optimal stopping set...

 $N = [I - Q]^{-1}$  is a fundamental matrix for a subst-c matrix Q.

# Optimal Stopping (OS) of Markov Chain (MC)

**T. Ferguson:** "Most problems of optimal stopping without some form of Markovian structure are essentially untractable...".

**OS Model**  $M = (X, P, c, g, \beta)$ :

- X finite (countable) state space,
- $P = \{p(x, y)\}$ , stochastic (transition) matrix
- c(x) one step cost function,
- g(x) terminal reward function,
- $\beta$  discount factor,  $0 \le \beta \le 1$
- (Z<sub>n</sub>) MC from a family of MCs defined by a Markov Model M = (X, P)
- $v(x) = \sup_{\tau \ge 0} E_x[\sum_{i=0}^{\tau-1} \beta^i c(Z_i) + \beta^{\tau} g(Z_{\tau})]$ value function

## Description of OS Continues

• **Remark !** absorbing state e, p(e, e) = 1,

 $p(x,y) \longrightarrow \beta p(x,y), \ p(x,e) = 1 - \beta,$ 

 $\beta \longrightarrow \beta(x) = P_x(Z_1 \neq e)$  probability of "survival".

•  $S_* = \{x : g(x) = v(x)\}$  optimal stopping set.

• 
$$Pf = Pf(x) = \sum_{y} p(x, y)f(y).$$

#### Theorem (Shiryayev 1969)

(a) The value function v(x) is the minimal solution of Bellman equation ...

 $v = \max(g, c + Pv),$ 

(b) if state space X is finite then set  $S_*$  is not empty and  $\tau_0 = \min\{n \ge 0 : Z_n \in S_*\}$  is an optimal stopping time. ...

## Basic methods of solving OS of MC, $c \equiv 0$

- The direct solution of the Bellman equation
- The value iteration method : one considers the sequence of functions  $v_n(x) = \sup_{0 \le \tau \le n} E_x \dots, v_{n+1}(x) = \max(g(x), Pv_n(x)),$  $v_0(x) = g(x)$ . Then  $v_0(x) \le v_1(x) \le \dots v_n(x)$  converges to v(x).
- The linear programming approach  $(|X| < \infty)$ ,  $\min \sum_{y \in X} v(y), v(x) \ge \sum_{y} p(x, y)v(y), v(x) \ge g(x), x \in X.$
- Davis and Karatzas (1994), interesting interpretation of the Doob-Meyer decomposition of the Snell's envelope
- The State Elimination Algorithm (SEA) Sonin (1995, 1999, 2005, 2008, 2010); A. Irle (1980, 2006);
  E. Presman (2009), continuos time

### State Elimination Algorithm for OS of MC

- OS = Bellman equation v(x) = max(g(x), c(x) + Pv(x));
- $M_1 = (X_1, P = P_1, c = c_1, g), S_* = S_{1*}$ . Three simple facts:
  - It may be *difficult* to find the states where it is optimal to stop,  $g(x) \ge c(x) + P_1v(x)$ , but it is easy to find a state (states) where it is optimal not to stop: do not stop if  $g(z)c(z) + P_1g(z) \le c(z) + P_1v(z)$ .
  - After identifying these states, set G, we can "eliminate" the subset D ⊂ G, and recalculate P<sub>1</sub> → P<sub>2</sub> and c<sub>1</sub> → c<sub>2</sub>, g. Elimination theorem: S<sub>1\*</sub> = S<sub>2\*</sub>, v<sub>1</sub> = v<sub>2</sub>. Repeat these steps until g(x) ≥ c<sub>k</sub>(x) + P<sub>k</sub>g(x) for all remaining x∈X<sub>k</sub>. Then
  - **Proposition 1.** Let M = (X, P, c, g) be an optimal stopping problem, and  $g(x) \ge c(x) + Pg(x)$  for all  $x \in X$ . Then X is the optimal stopping set in the problem M, and v(x) = g(x) for all  $x \in X$ .

## Eliminate state(s) z, (set D) and recalculate probabilities

Embedded Markov chain (Kolmogorov, Doeblin)  $M_1 = (X_1, P_1)$ ,  $D \subset X_1, X_2 = X_1 \setminus D = S$ ; MC  $(Z_n)$ ;  $\tau_0, \tau_1, ..., \tau_n, ...$ , the moments of zero, first, and so on, visits of  $(Z_n)$  to the set  $X_2$ . Let  $Y_n = Z_{\tau_n}, n = 0, 1, 2, ...$ 

#### Lemma (KD)

(a) The random sequence  $(Y_n)$  is a Markov chain in a model  $M_2 = (X_2, P_2)$ , where  $P_2 = \{p_2(x, y)\}$  is given by formula  $D \quad S$   $P_1 = \begin{bmatrix} Q_1 & T_1 \\ R_1 & P_{01} \end{bmatrix}$ ,  $P_2 = P_S = P_{01} + R_1 U_1 = P_{01} + R_1 N_1 T_1$ ,  $N_1 = N_D$  is a (transient) fundamental matrix, i.e.  $N_1 = (I - Q_1)^{-1}$ .  $N = I + Q + Q^2 + ... = (I - Q)^{-1}$ ,  $N = \{n(x, y)\}$ , U = NT.

### State Elimination Algorithm, $c \equiv 0$

If  $D = \{z\}$  then

$$p_2(x,y) = p_1(x,y) + p_1(x,z)n_1(z)p_1(z,y),$$

where  $n_1(z) = 1/(1 - p_1(z, z))$ . GTH/S algorithm (1985), inv. distr.

#### **State Elimination Algorithm**

### Forget about OS, just MCs.

Let  $M_i = (X_i, P_i)$  be two Markov models. Let i = 1, 2 and let  $h: X_1 \longrightarrow X_2$  be a mapping,  $H(t) = h^{-1}(t), t \in X_2$ .

Let  $(Z_n)$  be a MC in  $M_1$  and  $(Y_n)$  is defined by  $Y_n = h(Z_n)$ . Generally  $(Y_n)$  is not a MC. When it is ?

A necessary and sufficient condition for a MC to be "lumpable" (Kemeny, Snell, 1960), "mergeable" (Howard, 1971):

 $X_1$  is partitioned into  $H(t), t \in X_2$ ,

$$\sum_{y\in H(k)} p_1(x,y) = \sum_{y\in H(k)} p_1(x',y)$$

for any  $x, x' \in H(s)$  and any  $s, k \in X_2$ .

## Projections of Markov Models (MCs)

Model  $M_2$  is an *S*-projection of a basic model  $M_1$  (under *h*) (Seas.) if there is a stochastic matrix  $P_2$ ) and *m* dice, i.e.

 $p_1(x, y) = p_2(h(x), h(y))f(y|h(y))$ 

for all  $x, y \in X_1$ , where f(y|z) is a probability distribution on a set  $H(z) = h^{-1}(z) = \{y \in X_1 : h(y) = z\}$ , defined for each  $z \in X_2$ .

Model  $M_2$  is a *B*-projection of a basic model  $M_1$  (under *h*) if  $h(x) \neq h(y)$ , then transitions occur as above, and if h(x) = h(y), then

### $p_1(x, y) = p_2(h(x), h(x))q_1(x, y|h(x)),$

where stoch. matrix  $Q_1(k) = \{q_1(x|, y|k)\}$  is defined for each k. Model  $M_2$  is an *A*-projection of...if f(y|h(y)) is replaced by f(y|h(x), h(y)), i.e. there are not m but  $m^2$  dice, and maybe stochastic matrices  $Q_1(k), k \in X_2$ .

## Back to OS of Seasonal Observations

The key element in applying SEA is calculation of  $P_S$  using  $N_D$ , where...

The key element in the proof of Theorem 1 - the transformation of the equality  $P = U \times F$  into equality  $P_S = U_S \times F_S$ .

Let  $D \subset X_1$ ,  $S = X_1 \setminus D$  and we consider MC  $(Z_{n,D})$  stopped at  $S = X \setminus D$ .

By SEA to find matrix  $P_S$ , we have to find  $N_{1,D} = n_{1,D}(x, y), x, y \in X_1$ .

Let MC  $(U_{n,D})$  in model  $M_2$ , be the "projection" of MC  $(Z_{n,D})$ , defined by the equality  $U_{n,D} = h_D(Z_{n,D})$ , where function  $h_D(x) = h(x)$  if  $x \in D$  and  $h_D(X) = e$  if  $x \in S$ .

Let  $N_{2,D}$  be fundamental matrix for this MC.

Our goal is:  $P_S = U_S \times F_S$ . Not.  $F_d(A)$  diagonal matrix with elements  $F(A(k)) = \sum_{j \in A(k)} f(j|k)$ ,  $A(k) = A \cap H(k)$ , and  $U_A^F = UF_d(A)$ .

#### Theorem (2)

The fundamental matrices in the original and the projected models,  $N_{1,D}$  and  $N_{2,D}$  are related by the equalities valid for all  $x, y \in X_1$ ,  $n_{1,D}(x, y) = n_{2,D}(s, k)f(y|k)/F(D(k)), s = h(x), k = h(y)$ .

Using this theorem we can obtain the key lemma in PS 2010: Lemma

$$F_{S} = \{f_{S}(y|k) = f(y|k)/F(S(k))\},\$$

and

$$U_{S} = U_{S}^{F} + U_{D}^{F} (I - U_{D}^{F})^{-1} U_{S}^{F} = (I - U_{D}^{F})^{-1} U_{S}^{F}.$$

### **Open Problems.**

- Transformation of Fundamental matrices for all projections
- OS of Hidden MC
- explanation of world financial crisis
- Thank you for your attention ! Danke schon ! Merci beaucoup !

Spasibo !

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